PeerJ

Simphony: simulating large-scale, rhythmic data

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ABSTRACT

Simulated data are invaluable for assessing a computational method's ability to distinguish signal from noise. Although many biological systems show rhythmicity, there is no general-purpose tool to simulate large-scale, rhythmic data. Here we present Simphony, an R package for simulating data from experiments in which the abundances of rhythmic and non-rhythmic features (e.g., genes) are measured at multiple time points in multiple conditions. Simphony has parameters for specifying experimental design and each feature's rhythmic properties (e.g., amplitude and phase). In addition, Simphony can sample measurements from Gaussian and negative binomial distributions, the latter of which approximates read counts from RNA-seq data. We show an example of using Simphony to evaluate the accuracy of rhythm detection. Our results suggest that Simphony will aid experimental design and computational method development. Simphony is thoroughly documented and freely available at https://github.com/hugheylab/simphony.

Subjects Bioinformatics, Computational Biology, Data Science **Keywords** Rhythms, Circadian, Simulation, Gene expression, Transcriptome

INTRODUCTION

Rhythms are ubiquitous across domains of life and across timescales, from hourly division of bacteria (*Cooper & Helmstetter*, 1968) to seasonal growth of trees (*Kramer*, 1936). These biological rhythms are often driven by systems of genes and proteins. Prominent examples are the systems underlying circadian rhythms, which have a period of approximately 24 h and have been observed in species across the biosphere (*Young & Kay*, 2001) and throughout the body of multicellular organisms (*Yoo et al.*, 2004; *Zhang et al.*, 2014).

To interrogate these rhythmic biological systems, researchers are increasingly using technologies that measure the abundance of thousands of molecules in parallel (e.g., quantifying the transcriptome by RNA-Seq). The critical decisions then become how to design the experiments and how to analyze the data. For example, there are now numerous methods for detecting rhythms in high-dimensional data (*Yang & Su, 2010*; *Hughes, Hogenesch & Kornacker, 2010*; *Thaben & Westermark, 2014*; *Wu et al., 2016*). A valuable aid to such decisions is simulation. In simulated data, unlike in experimental data, the ground truth is known (e.g., whether a gene is rhythmic). Consequently, to the extent that simulated data possess the essential features of experimental data, simulation can be used to estimate a method's ability to distinguish signal from noise (*Deckard et al., 2013*).

Submitted 19 December 2018 Accepted 15 April 2019 Published 23 May 2019

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Academic editor Andrew Gray

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DOI 10.7717/peerj.6985

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Simulated data are also typically faster and less expensive to generate than experimental data, especially omics data from high-resolution time courses.

Unfortunately, there is a shortage of publicly available tools for simulating rhythmic data. This forces researchers to create their own simulations from scratch (*Deckard et al.*, 2013; Singer & Hughey, 2018) or to forgo simulations altogether. Although several tools exist to simulate particular types of transcriptome data (Dembélé, 2013; Frazee et al., 2015; Zappia, Phipson & Oshlack, 2017), most are not designed to simulate data from time-course experiments. One exception is Polyester (Frazee et al., 2015), which can simulate RNA-seq reads from multiple time points and conditions. However, Polyester models many aspects of the sequencing process, which incurs a computational burden and may not be directly relevant for designing experiments to collect rhythmic data or evaluating methods to analyze such data. Recognizing a gap, Hughes et al. (2017) recently developed CircaInSilico, a web-based application for simulating rhythmic data. Although CircaInSilico has a convenient user interface, it has several limitations-for example, the simulated rhythms can only be sinusoidal. In addition, even though read counts from RNA-seq data are often modeled using a negative binomial distribution (Robinson & Smyth, 2007), CircaInSilico can only simulate Gaussian noise. Thus, there is still a need for a flexible tool to simulate large-scale, rhythmic data.

To address this need, we developed a simulation package called Simphony. Simphony has adjustable parameters for specifying experimental design and modeling rhythms, including the ability to sample from Gaussian and negative binomial distributions. Simphony is implemented in R, thoroughly documented, and freely available at https://github.com/hugheylab/simphony.

MATERIALS AND METHODS

Simulating rhythmic data using Simphony

Simphony simulates experiments in which the abundances of rhythmic and non-rhythmic features (e.g., genes) are measured at multiple time points in one or more conditions (Table 1). Within a given simulated experiment (i.e., a simulation), the expected abundance m of feature i in condition k at time t is modeled as

 $m_{ik}(t) = a_{ik}(t) \cdot f_{ik}(\frac{2\pi}{\tau_{ik}} \cdot (t + \phi_{ik})) + b_{ik}(t),$

where *a* is the amplitude, *f* is a periodic function with period 2π (by default, $f(\theta) = sin(\theta)$), τ is the period of rhythmic changes in abundance (by default, 24), ϕ is the phase, and *b* is the baseline abundance. If *a* and *b* are constant and $f(\theta) = sin(\theta)$, the model is equivalent to cosinor. Time-dependent *a* can create damped rhythms, whereas time-dependent *b* can create drift. Non-rhythmicity is defined by a = 0.

Given $m_{ik}(t)$, Simphony samples measurements from one of two families of distributions: Gaussian and negative binomial. The former represents an idealized experimental scenario, whereas the latter approximates read counts from RNA-seq. For Gaussian sampling, the abundance of feature *i* in sample *j* belonging to condition *k* follows $Y_{ii} \sim N(m_{ik}(t_i), \sigma_{ik}^2)$,

Table 1Available options in Simphony.	
Type of parameter	Parameters
Experimental design	Time points:
	• First and last time points, interval, and number of samples per time point
	• Specified time points and number of samples per time point
	• Drawn from a uniform distribution, first and last possible time points, and total number of samples
	Number of conditions
	Number of features in each group
Properties of abundance	Rhythmic shape
(per feature group per condition)	Period
	Phase
	Amplitude (can be time-dependent)
	Baseline (can be time-dependent)
Sampling measurements	Family:
	• Gaussian
	• Negative binomial
	Standard deviation (if Gaussian)
	Mean-dispersion relationship (if negative binomial)

where σ^2 is the variance (by default, 1). For negative binomial sampling, we follow a similar strategy to DESeq2 (*Love, Huber & Anders, 2014*) and Polyester (*Frazee et al., 2015*), such that

$$Y_{ij} \sim NB(\mu = 2^{m_{ik}(t_j)}, \alpha = g_{ik}(2^{m_{ik}(t_j)}))$$

where μ is the expected counts, α is the dispersion (the variance of a negative binomial distribution is $Var(Y) = \mu + \alpha \mu^2$), and g is a function that maps expected counts to dispersion. The default g was estimated from RNA-seq data from mouse liver (see the next section for details).

Experimental design in Simphony is specified in one of three ways: (1) first and last time points, interval between time points, and number of samples per time point per condition, (2) exact time points and number of samples per time point per condition, or (3) time points sampled from a uniform distribution, range of possible time points, and total number of samples per condition. By default, Simphony uses option (1), with first and last time points of 0 and 48, interval between time points of 3, and number of samples per time point of 2.

The Simphony R package has two dependencies: data.table (*Dowle & Srinivasan, 2018*) and foreach (*Calaway, Microsoft & Weston, 2017*).

Estimating statistical properties of experimental RNA-seq data

To estimate the relationship between expected counts and dispersion in real RNA-seq data, we used PRJNA297287 (*Atger et al., 2015*). We used the samples that were collected in quadruplicate from livers of wild-type, ad libitum-fed mice every 2 h for 24 h in LD

12:12 (48 samples total). We downloaded the raw reads, then quantified gene-level counts using Salmon v0.11.3 (*Patro et al., 2017*) and tximport v1.8.0 (*Soneson, Love & Robinson, 2015*). We kept the 15,069 genes that had at least 10 counts in half of the samples. We used DESeq2 v1.20.0 to estimate parametric and local regression-based mean-dispersion curves (*Love, Huber & Anders, 2014*) (Fig. S1A). The input to DESeq2 included a design matrix based on cosinor regression, so that dispersion estimates were not biased by variation in expression due to a daily rhythm. Compared to the parametric mean-dispersion curve, the local regression-based curve had a considerably lower root-mean-squared error (0.94 compared to 1.09, in units of log dispersion), so we set it as the default in Simphony (*g* in the equation above). DESeq2 also provided an estimate of the variance of the residual log dispersion (around the curve). Finally, we used fitdistrplus v1.0-14 (*Delignette-Muller & Dutang, 2015*) to approximate the distribution of mean normalized counts as log-normal. The Simphony documentation includes an example of how to sample from the estimated distributions of residual log dispersion and mean normalized counts (Fig. S1B).

Validating statistical properties of simulated data

We performed multiple simulations to validate the statistical properties of data generated by Simphony. Each simulation had time points spaced 0.1 h apart (period of 24 h), 100 samples per time point, and one feature for each combination of parameter values related to measurements. Simulations based on negative binomial sampling used the default function for calculating dispersion.

To validate mean and standard deviation, we simulated non-rhythmic abundance (amplitude of 0) based on Gaussian and negative binomial sampling. For the simulation using Gaussian sampling, we varied the desired mean and standard deviation. For the simulation using negative binomial sampling, we varied the desired mean \log_2 counts. In both cases, we then calculated the empirical mean and standard deviation (Table S1).

To validate amplitude and phase, we simulated rhythmic abundance based on Gaussian and negative binomial sampling (using the default $f(\theta) = sin(\theta)$). For both types of sampling, we varied the desired amplitude and phase. For the simulation based on Gaussian sampling, we used the limma R package v3.38.3 (*Smyth*, 2004; *Ritchie et al.*, 2015) to fit each feature's abundance to a linear model that had terms for $cos(\frac{2\pi}{\tau}t)$ and $sin(\frac{2\pi}{\tau}t)$ (cosinor regression). We then used the model coefficients to estimate each feature's amplitude and phase according to the trigonometric identity $a \cdot cos\theta + b \cdot sin\theta = c \cdot sin(\theta + \phi)$, where $c = \sqrt{a^2 + b^2}$ and $\phi = \frac{\pi}{2} - atan2(b, a)$ (Table S2). For the simulation based on negative binomial sampling, we followed a similar procedure, except we log-transformed the counts before passing them to limma.

Detecting rhythmicity in simulated data

We calculated gene-wise p-values of rhythmicity using JTK_CYCLE v3.1 (*Hughes*, *Hogenesch & Kornacker*, 2010) after transforming the expression values, sampled from the negative binomial family, using log₂(counts+1). We used the p-values and the precrec R package v0.9.1 (*Saito & Rehmsmeier*, 2017) to calculate the area under the receiver operating characteristic (ROC) curve for distinguishing non-rhythmic genes from each group of rhythmic genes (specified by rhythm amplitude and baseline in log₂ counts).

RESULTS

To validate the statistical properties of data generated by Simphony, we simulated data covering a range of parameter values for the Gaussian and negative binomial families. To ensure that the properties approached their asymptotic values, time points were spaced 0.1 h apart (period of 24 h), each with 100 samples. For non-rhythmic abundance, we verified that the observed mean and standard deviation corresponded to the expected values (Table S1). For rhythmic abundance, we verified that the observed amplitude and phase corresponded to the expected values (Table S2).

To highlight Simphony's flexibility, we simulated gene expression from a variety of patterns, including ones in which the rhythm amplitude or baseline expression was time-dependent (i.e., non-stationary). For each pattern, we sampled expression values from the Gaussian and negative binomial families (Fig. 1). These patterns are only examples— Simphony can simulate data from any rhythmic waveform or non-stationary trend provided as a function in R. We also simulated an experiment in which each of 200 genes had a different rhythm amplitude and phase (Fig. 2), and an experiment having two conditions, in which genes' rhythms had a different amplitude, phase, or period in each condition (Fig. S2).

To show an example of Simphony's utility, we created simulations to quantify how the accuracy of rhythm detection depends on experimental and biological parameters. We simulated experiments having various intervals between time points and one sample per time point. Each simulation included 20,000 genes spanning a range of values for baseline expression and rhythm amplitude (including amplitude 0 for non-rhythmic genes) (Fig. S3A). Because Simphony is not designed to detect rhythmicity, we calculated each gene's *p*-value of rhythmicity in each simulation using JTK_CYCLE, then calculated the area under the ROC curve for distinguishing non-rhythmic genes from each group of rhythmic genes. As expected, rhythm detection improved as rhythm amplitude increased or the interval between time points decreased (Fig. 3A). Rhythm detection also improved as baseline expression increased (and thus as the standard deviation of log-transformed counts of non-rhythmic genes decreased; Fig. 3B and Fig. S3B).

DISCUSSION

Simphony is a versatile framework for simulating rhythmic data. Although Simphony is especially apt for simulating transcriptome data, it is general enough to simulate data of various types (e.g., bioluminescence). A future objective is to use simulated data from Simphony to comprehensively benchmark computational methods for detecting rhythmicity. Simphony's flexibility will be key to mimicking the diversity of rhythms seen in practice. Simphony's ability to simulate non-stationary trends in particular is critical, since the possibility of non-stationarity is one reason the guidelines for genome-scale analysis of biological rhythms recommend collecting data from at least two cycles (*Hughes*)





et al., 2017). Ultimately, we anticipate that Simphony will inform the design of experiments for interrogating rhythmic biological systems and the development of methods for analyzing data containing rhythmic signals.



Figure 2 Example of medium-scale simulation in Simphony. Expression values of 200 rhythmic genes, each gene with its own amplitude and phase, were sampled from the Gaussian family. Rhythms followed a sinusoid. Each row in the heatmap corresponds to a gene, each column to a time point. For ease of visualization, sampled expression values were cropped to be between -4 and 4.

Full-size DOI: 10.7717/peerj.6985/fig-2



Figure 3 Example of evaluating rhythm detection using data generated by Simphony. Simulations had various values of the interval between time points and one replicate per time point. Each simulation included 20,000 genes having various values of baseline expression and rhythm amplitude (including amplitude 0). Rhythms followed a sinusoid of period 24 h. Expression values were sampled from the negative binomial family. Gene-wise p-values of rhythmicity from JTK_CYCLE were used to calculate the area under the ROC curve (AUROC) for distinguishing non-rhythmic genes from each group of rhythmic genes. (A) AUROC vs. rhythm amplitude and interval, for genes with a baseline log2 counts of 8. (B) AUROC vs. rhythm amplitude and baseline expression, for the simulation with an interval of 2 h. AUROC of 0.5 corresponds to random detection, while AUROC of 1 corresponds to perfect detection.

Full-size DOI: 10.7717/peerj.6985/fig-3

ADDITIONAL INFORMATION AND DECLARATIONS

Funding

This work was supported by NIH R35GM124685 to Jacob J. Hughey. The funders had no role in study design, data collection and analysis, decision to publish, or preparation of the manuscript.

Grant Disclosures

The following grant information was disclosed by the authors: NIH: R35GM124685.

Competing Interests

The authors declare there are no competing interests.

Author Contributions

• Jordan M. Singer, Darwin Y. Fu and Jacob J. Hughey conceived and designed the experiments, performed the experiments, analyzed the data, prepared figures and/or tables, authored or reviewed drafts of the paper, approved the final draft.

Data Availability

The following information was supplied regarding data availability:

The simphony R package is available on GitHub: https://github.com/hugheylab/ simphony. Data, code, and results for this study are available on Figshare: https: //doi.org/10.6084/m9.figshare.7441355.

Supplemental Information

Supplemental information for this article can be found online at http://dx.doi.org/10.7717/ peerj.6985#supplemental-information.

REFERENCES

- Atger F, Gobet C, Marquis J, Martin E, Wang J, Weger B, Lefebvre G, Descombes P, Naef F, Gachon F. 2015. Circadian and feeding rhythms differentially affect rhythmic mRNA transcription and translation in mouse liver. *Proceedings of the National Academy of Sciences of the United States of America* 112:E6579–E6588 DOI 10.1073/pnas.1515308112.
- **Calaway R, Microsoft, Weston S. 2017.** foreach: provides foreach looping construct for R. Comprehensive R Archive Network (CRAN). *Available at https://cran.r-project.org/package=foreach*.
- Cooper S, Helmstetter CE. 1968. Chromosome replication and the division cycle of Escherichia coli B/r. *Journal of Molecular Biology* **31**:519–540 DOI 10.1016/0022-2836(68)90425-7.
- Deckard A, Anafi RC, Hogenesch JB, Haase SB, Harer J. 2013. Design and analysis of large-scale biological rhythm studies: a comparison of algorithms for

detecting periodic signals in biological data. *Bioinformatics* **29**:3174–3180 DOI 10.1093/bioinformatics/btt541.

- **Delignette-Muller M, Dutang C. 2015.** fitdistrplus: an R Package for fitting distributions. *Journal of Statistical Software, Articles* **64**:1–34 DOI 10.18637/jss.v064.i04.
- **Dembélé D. 2013.** A flexible microarray data simulation model. *Microarrays* **2**:115–130 DOI 10.3390/microarrays2020115.
- **Dowle M, Srinivasan A. 2018.** data.table: extension of 'data.frame'. Comprehensive R Archive Network (CRAN).
- **Frazee AC, Jaffe AE, Langmead B, Leek JT. 2015.** Polyester: simulating RNA-seq datasets with differential transcript expression. *Bioinformatics* **31**:2778–2784 DOI 10.1093/bioinformatics/btv272.
- Hughes ME, Abruzzi KC, Allada R, Anafi R, Arpat AB, Asher G, Baldi P, De Bekker C, Bell-Pedersen D, Blau J, Brown S, Ceriani MF, Chen Z, Chiu JC, Cox J, Crowell AM, DeBruyne JP, Dijk D-J, DiTacchio L, Doyle FJ, Duffield GE, Dunlap JC, Eckel-Mahan K, Esser KA, FitzGerald GA, Forger DB, Francey LJ, Fu Y-H, Gachon F, Gatfield D, De Goede P, Golden SS, Green C, Harer J, Harmer S, Haspel J, Hastings MH, Herzel H, Herzog ED, Hoffmann C, Hong C, Hughey JJ, Hurley JM, De la Iglesia HO, Johnson C, Kay SA, Koike N, Kornacker K, Kramer A, Lamia K, Leise T, Lewis SA, Li J, Li X, Liu AC, Loros JJ, Martino TA, Menet JS, Merrow M, Millar AJ, Mockler T, Naef F, Nagoshi E, Nitabach MN, Olmedo M, Nusinow DA, Ptáček LJ, Rand D, Reddy AB, Robles MS, Roenneberg T, Rosbash M, Ruben MD, Rund SSC, Sancar A, Sassone-Corsi P, Sehgal A, Sherrill-Mix S, Skene DJ, Storch K-F, Takahashi JS, Ueda HR, Wang H, Weitz C, Westermark PO, Wijnen H, Xu Y, Wu G, Yoo S-H, Young M, Zhang EE, Zielinski T, Hogenesch JB. 2017. Guidelines for genome-scale analysis of biological rhythms. *Journal of Biological Rhythms* 32:380–393 DOI 10.1177/0748730417728663.
- Hughes ME, Hogenesch JB, Kornacker K. 2010. JTK_CYCLE: an efficient nonparametric algorithm for detecting rhythmic components in genome-scale data sets. *Journal of Biological Rhythms* 25:372–380 DOI 10.1177/0748730410379711.
- **Kramer PJ. 1936.** Effect of variation in length of day on growth and dormancy of trees. *Plant Physiology* **11**:127–137 DOI 10.1104/pp.11.1.127.
- Love MI, Huber W, Anders S. 2014. Moderated estimation of fold change and dispersion for RNA-seq data with DESeq2. *Genome Biology* 15:Article 550 DOI 10.1186/s13059-014-0550-8.
- Patro R, Duggal G, Love MI, Irizarry RA, Kingsford C. 2017. Salmon provides fast and bias-aware quantification of transcript expression. *Nature Methods* 14:417–419 DOI 10.1038/nmeth.4197.
- Ritchie ME, Phipson B, Wu D, Hu Y, Law CW, Shi W, Smyth GK. 2015. limma powers differential expression analyses for RNA-sequencing and microarray studies. *Nucleic Acids Research* **43**:e47 DOI 10.1093/nar/gkv007.
- Robinson MD, Smyth GK. 2007. Moderated statistical tests for assessing differences in tag abundance. *Bioinformatics* 23:2881–2887 DOI 10.1093/bioinformatics/btm453.

- Saito T, Rehmsmeier M. 2017. Precrec: fast and accurate precision—recall and ROC curve calculations in R. *Bioinformatics* 33:145–147 DOI 10.1093/bioinformatics/btw570.
- **Singer JM, Hughey JJ. 2018.** LimoRhyde: a flexible approach for differential analysis of rhythmic transcriptome data. *Journal of Biological Rhythms* 748730418813785 DOI 10.1177/0748730418813785.
- **Smyth GK. 2004.** Linear models and empirical bayes methods for assessing differential expression in microarray experiments. *Statistical Applications in Genetics and Molecular Biology* **3**:Article 3 DOI 10.2202/1544-6115.1027.
- Soneson C, Love MI, Robinson MD. 2015. Differential analyses for RNA-seq: transcriptlevel estimates improve gene-level inferences. *F1000 Research* 4:Article 1521 DOI 10.12688/f1000research.7563.1.
- Thaben PF, Westermark PO. 2014. Detecting rhythms in time series with RAIN. *Journal of Biological Rhythms* 29:391–400 DOI 10.1177/0748730414553029.
- Wu G, Anafi RC, Hughes ME, Kornacker K, Hogenesch JB. 2016. MetaCycle: an integrated R package to evaluate periodicity in large scale data. *Bioinformatics* 32:3351–3353 DOI 10.1093/bioinformatics/btw405.
- Yang R, Su Z. 2010. Analyzing circadian expression data by harmonic regression based on autoregressive spectral estimation. *Bioinformatics* 26:i168–74 DOI 10.1093/bioinformatics/btq189.
- Yoo S-H, Yamazaki S, Lowrey PL, Shimomura K, Ko CH, Buhr ED, Siepka SM, Hong H-K, Oh WJ, Yoo OJ, Menaker M, Takahashi JS. 2004. PERIOD2::LUCIFERASE real-time reporting of circadian dynamics reveals persistent circadian oscillations in mouse peripheral tissues. *Proceedings of the National Academy of Sciences of the United States of America* 101:5339–5346 DOI 10.1073/pnas.0308709101.
- **Young MW, Kay SA. 2001.** Time zones: a comparative genetics of circadian clocks. *Nature reviews. Genetics* **2**:702–715 DOI 10.1038/35088576.
- Zappia L, Phipson B, Oshlack A. 2017. Splatter: simulation of single-cell RNA sequencing data. *Genome Biology* 18:Article 174 DOI 10.1186/s13059-017-1305-0.
- Zhang R, Lahens NF, Ballance HI, Hughes ME, Hogenesch JB. 2014. A circadian gene expression atlas in mammals: implications for biology and medicine. *Proceedings of the National Academy of Sciences of the United States of America* 111:16219–16224 DOI 10.1073/pnas.1408886111.